

#### **REVIEW ARTICLE**



# The rise to power of the microbiome: power and sample size calculation for microbiome studies

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A priori power and sample size calculations are crucial to appropriately test null hypotheses and obtain valid conclusions from all clinical studies. Statistical tests to evaluate hypotheses in microbiome studies need to consider intrinsic features of microbiome datasets that do not apply to classic sample size calculation. In this review, we summarize statistical approaches to calculate sample sizes for typical microbiome study scenarios, including those that hypothesize microbiome features to be the outcome, the exposure or the mediator, and provide relevant R scripts to conduct some of these calculations. This review is intended to be a resource to facilitate the conduct of sample size calculations that are based on testable hypotheses across several dimensions of the microbiome. Implementation of these methods will improve the quality of human or animal microbiome studies, enabling reliable conclusions that will generalize beyond the study sample.

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#### INTRODUCTION

The human gut is host to a community of microbes (bacteria, archaea, fungi, viruses and phages) referred to as the gut microbiota, with their combined functions known as the gut microbiome<sup>1</sup>. Since the start of the Human Microbiome Project (HMP), wide-ranging, genome-scale community research in 2008, large microbiome studies have aimed to characterize the genetic diversity of microbial populations living in and on humans. This is being achieved by applying next generation sequencing technology and exploring the diversity and composition of these microbial communities in the context of human body functions and mechanisms that lead to diseases<sup>2</sup>. This promising field of research may contribute to the prognosis of clinical outcomes through microbial biomarkers—any measurement allowing an intercommunication between a biological system and a potential risk, which may be chemical, physical, or biological<sup>3</sup>. There is also a growing interest in the influence of the microbiome on human

As with any research, clear and testable research hypotheses are required to conduct high quality studies of the microbiome. To improve the quality and consistency of microbiomics research reporting, guidelines have recently been published to critically appraise microbiome studies, which include sample size or power calculation as a criterion for a well-conducted study<sup>4</sup>. Most human microbiome studies aim to identify the relationship between microbiome features and a biological or clinical condition, an environmental exposure or medical intervention. Since much is unknown about the microbiome, and the datasets are vast and

often not normally distributed, a variety of data driven techniques have been developed<sup>5</sup>. However, with the accumulation of microbiome evidence, hypothesis-based comparisons are increasingly possible.

Microbiome data arise from sequencing a marker gene such as the 16S ribosomal RNA gene for bacteria, and the ITS marker for fungi, or from metagenomic sequencing of the entire DNA within a community. The sequencing data are then summarized into a series of counts. These counts may represent Amplicon Sequence Variants (ASVs) capturing single nucleotide differences between sequences, or they may be clustered results such as the counts of unique Operational Taxonomic Units (OTUs), or taxa abundances<sup>6</sup>. Therefore, sample size calculations may focus on the count of one single cluster, taxon, ASV, or they may be based on the full spectrum of counts. For research questions based on single counts or abundances, sample size calculations can use standard formulae applicable in many domains. However, when aiming to determine if a difference exists in the whole spectrum of counts or the whole microbial community, microbiome-specific methods are required. For example, it is non-trivial to implement sample size calculations for beta-diversity measures<sup>7</sup>, since realistic distance matrices are needed between pairs of samples and between groups of samples. Either appropriate pilot data must be found, or a complex simulation study must be undertaken.

On behalf of IMPACTT (Integrated Microbiome Platforms for Advancing Causation Testing and Translation), we have compiled a comprehensive guide for sample size calculation for microbiome studies. After a brief introduction to the general concepts behind

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**Table 1.** Definition of selected statistical concepts.

Concept	Definition
Type I error (false positive rate)	The incorrect rejection of a true null hypothesis. If there is one true null hypothesis ( $H_0$ ), the probability of rejecting it is usually denoted $a$ .
Type II error (false negative rate)	The failure to reject a false null hypothesis. If the null hypothesis is false ( $H_A$ ), the probability of not rejecting the null hypothesis is usually denoted $\beta$ .
Power of the test	The probability of rejecting the null hypothesis when, in fact, it is false. Using the notation above, this is $1 - \beta$ .
Effect size	A quantitative measure of association related to the proposed statistical test to be used. When there are two groups, such as a control group and an experimental group, the desired measure of association is often the difference between-group means or medians. It usually refers to a standardized difference, where the expected difference between the measure of interest is divided by an estimate of the standard deviation of this measure.

sample size calculation, we first set the stage for microbiome sample size calculations through our decision tree lens. Then, we provide examples of sample size calculations for each node in the decision tree.

### SAMPLE SIZE AND STATISTICAL POWER CALCULATIONS: GENERAL CONSIDERATIONS

Sample size and statistical power calculations should be performed at the design stage of a microbiome study. Ensuring good power means that the study conclusions, relating to an association or an effect of interest, are likely to be valid, and that the conclusions will generalize beyond the study sample to future similar studies. Such calculations depend on three concepts, the required sample size (*N*), the researcher's tolerance for making errors in conclusions (values of Type I or II error)<sup>8,9</sup>, and the magnitude of the effect or association that one is trying to detect<sup>10,11</sup>.

#### Type I and Type II errors

Researchers must decide in advance what levels of error they will tolerate. Two types of errors can be anticipated: type I error and type II error, defined in Table 1. Sample size calculations depend on a statistic or a measure of association, and on the behavior of this statistic when there is no association, i.e., under the null hypothesis. Following the notation in the previous table, suppose one performed a t-test to compare means in two groups. One would reject the null hypothesis of no difference between twogroup means if  $p < a^*$ , with  $a^*$  being the significance for rejecting a null hypothesis determined before starting the experiment. When one performs a single test of hypothesis, then  $a^*$  is the same as the probability of making a type I error, that is  $\alpha^* = \alpha$ . In most studies, a type I error of 1 or 5% is a common choice<sup>9</sup>. The main aim of this parameter is to control the probability of making false positive conclusions. When many tests of hypothesis are performed, the threshold for rejection of the null will need to be smaller than the desired overall type I error for the entire study, so  $\alpha^* < \alpha^{12}$ .

A power calculation assists in choosing the number of subjects needed to prevent a type II error, denoted  $\beta$ ; power is defined as one minus the type II error or  $1-\beta$ . The choice of an appropriate value for the type II error,  $\beta$ , can be quite context dependent. For example, in clinical studies a commonly used threshold is 20%, which indicates an 80% chance of finding a true association. This may be appropriate to avoid misuse of resources, since it is usually necessary to increase sample size to increase power<sup>9</sup>. However, there are situations where making a type II error would be an unfortunate study outcome, and where smaller values of type II errors may be desired<sup>8</sup>. For example, when there are no effective approved treatments for a clinical condition, it would be important not to miss the potential benefit of a new treatment.

### Effect size

The effect size describes the magnitude of the difference of interest. For instance, if the species richness has a median of 32 for the control group and 15 for a treatment group exposed to

antibiotics, one might define an effect size as the differences between the medians, i.e., 17.

The strength of the association between two variables depends not only on the size of the difference—17 in the previous example —but also on the variability of this quantity. In the example above, one would need the variance of species richness' measures across samples from the same treatment group. Such variance estimates can often be obtained from previous studies applying similar methodology and using similar measures, or from a pilot study.

The definition of effect size depends on the study design <sup>13</sup> (i) to compare a continuous measure, such as species richness between two groups, the effect size will normally be a standardized mean difference, as described above; (ii) to compare the presence or absence of one particular species between two groups, the effect size may be an odds ratio or difference in proportions; (iii) to determine whether microbial community composition or alphadiversity is associated with a continuous measure, the effect size might be based on a Pearson correlation (r). Long-established standards define  $r \sim 0.1$  as a small effect size,  $r \sim 0.3$  as medium, and  $r \geq 0.5$  as large <sup>13</sup>.

In microbiome studies, the number of communities present and the community structures greatly vary across different study designs and platforms, even within a single site of sampling such as the human gut. Therefore, obtaining accurate estimates of variability can be challenging. When possible, examining effect sizes in a study of similar design is the best option for obtaining realistic estimates. However, when such studies are not available, it may be necessary to look for previous studies that are as close as possible in design and goals to find appropriate variance estimates. In general, larger sample sizes are needed when one desires smaller type I error, smaller type II error, and/or smaller effect sizes 14.

### **BASIC FORMULAE FOR SAMPLE SIZE CALCULATION**

We have provided the most commonly used formulae for sample size calculations in Table 2.

### WORKFLOW FOR SAMPLE SIZE CALCULATIONS IN MICROBIOME STUDIES

Figure 1 lays out a decision tree classification system for commonly used microbiome study designs. To help with choosing an appropriate sample size or power calculation method, each node number in Fig. 1 refers to a subsection of "Sample size calculations associated with each node of the decision tree", that provides suggestions for sample size calculation approaches with several worked examples.

### SAMPLE SIZE CALCULATIONS ASSOCIATED WITH EACH NODE OF THE WORKFLOW CHART

In this section, following the workflow chart in Fig. 1 from top to bottom, we provide specific references and worked examples of sample size calculations. Each subsection corresponds to one node in Fig. 1; and one row in Table 3 shows key formulae or

**Table 2.** Standard formulae for simple sample size calculations.

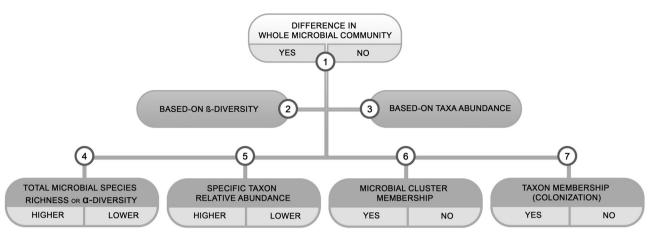
Type of sample size calculations	Standard formulae	
Comparison of two means in normally distributed continuous data	If the two groups contain the same number of samples, the required sample size per group can be calculated as $n = \frac{2(Z_{1-\frac{a}{2}} + Z_{1-\beta})^2}{\alpha^2}$ , Equation A	
	where:	
	• $n =$ the required sample size in each group	
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	• $\Delta = \frac{\mu_1 - \mu_2}{\sigma}$ , is the effect size, where $\mu_1$ and $\mu_2$ are two populations means and $\sigma$ is the common variance, $\sigma_1 = \sigma_2 = \sigma$	
	• $Z_{1-\alpha}, Z_{1-\beta}$ are the upper tail normal quantiles associated with the desired type I and type II errors, and $\beta$ , respectively.	
	If the two groups are not equally sized, then let parameter $r$ denote the ratio of the number of individuals in the larger group divided by the number of individuals in the smaller group. The sample size of the smaller group for a two-sided $Z$ test is given as follows:	
	$n_2 = \frac{r+1}{r} \frac{(Z_{1-\frac{\alpha}{2}} + Z_{1-\beta})^2}{\Delta^2},$	
	and the sample size for the larger group,	
	$n_1 = r n_2$ , Equation B	
Comparison of the difference in proportions between two groups	The following formula is used to estimate per group sample size for a difference in proportions, assuming equal sample sizes in both groups <sup>44</sup> $n = \frac{(Z_{1-\frac{q}{2}} + Z_{1-\beta})^2 (P_1(1-P_1) + P_2(1-P_2))}{(P_1-P_2)^2}, \text{Equation C}$	
	where:	
	• $P_1$ = the proportion in the first group	
	• $P_2$ = the proportion in the second group.	
	• $Z_{1-\frac{\alpha}{2}} = 1.96 \ (\alpha = 0.05), Z_{1-\beta} = 0.84 \ (\beta = 0.20)$	
	• $P_1 - P_2 =$ Effect Size (difference in proportions).	
	If $n_1 \neq n_2$ , the ratio between the sample sizes of the two groups is $r = \frac{n_1}{n_2}$ . Then the formulas that are used to compute sample size and power <sup>43</sup> are given below, respectively:	
	$n_1 = r n_2$ , and	
	$n_2 = (\frac{p_1(1-p_1)}{r} + p_2(1-p_2))(\frac{z_1 - \frac{a}{2} + z_1 - \beta}{p_1 - p_2})^2$ , Equation D	
	For a test statistic $t$ , and for the normal density function $\phi(.)$ , power can be estimated by:	
	$1 - \beta = \phi(t - Z_{1-\frac{q}{2}}) + \phi(-t - Z_{1-\frac{q}{2}})$ , Equation E	
Comparison of the odds between two groups	The following formula is used to estimate per group sample size for <i>an odds ratio</i> , assuming equal sample sizes in both groups <sup>45</sup> :	
	• Define $\kappa = \frac{n_1}{n_2}$ as the ratio of the numbers of individuals in the groups, 1 and 2, where 1 and 2 are defined based on the exposure variable X.	
	Define the odds ratio (OR) as:	
	$OR = \frac{p_1(1-p_2)}{p_2(1-p_1)},$	
	where $p_1$ and $p_2$ are proportions of the samples where the taxon abundance is above the chosen threshold (e.g., median) in the two exposure groups.	
	Then:	
	$n_1 = \kappa n_2$ , and	
	$n_2 = (\frac{1}{\kappa p_1(1-p_1)} + \frac{1}{p_2(1-p_2)})(\frac{z_{(a/2)} + z_{1-\beta}}{\ln(OR)})^2$ , Equation F	
Sample size based on correlations	The sample size required to test the hypothesis that the population correlation $(\rho_{yx})$ is equal to a specified value $(h;$ usually we set $h=0$ and test $\rho_{yx}=0$ ) for a given confidence level $(1-a)$ and power $(1-\beta)$ is approximately:	
	$n=3+((z_{(1-\frac{\alpha}{2})}+z_{1-\beta})^2/(\widetilde{ ho^*}_{yx}-h^*)^2),$ Equation G	
	where $\tilde{\rho}^*_{yx} = \ln(\frac{1+\tilde{\rho}_{yx}}{1-\tilde{\rho}_{yx}})/2$ is called the Fisher transformation of $\tilde{\rho}_{yx}$ , the planning value for $\rho_{yx}$ . The desired null hypothesis value $h$ must also be transformed with the Fisher transformation to $h^*$ in Equation G, and the numerator captures the adjustment necessary to obtain the desired type I and type II errors <sup>46</sup> .	

references for sample size and power calculations for the specific hypothesis being tested.

# Comparing microbial community structure between groups versus within groups using beta-diversity or distance metrics (Fig. 1, node 2)

The most commonly used analytic approach when working with the full spectrum of microbiome counts is to use beta-

diversity, or measures of distance or dissimilarity between samples. To estimate sample size or power, one must choose first a distance metric, then find or generate plausible distances that are relevant for the proposed study, that is, obtain likely distances between pairs of samples. Sample size calculations are then based on the distributions of these distances, by comparing distances for pairs from the same group to pairs from different groups. The first row of Table 3 (Comparing microbial community



**Fig. 1 Workflow for hypothesis-specific samples size equations.** An overview of the workflow to determine sample size equation for specific hypothesis. Each node is numbered and described further in section. Sample size calculations associated with each node of the workflow chart. Node 1, at the top of the workflow, describes a conceptual exploration of microbiome patterns, where a non-specific characterization of the full microbiome diversity, between and within samples, is of interest. This might imply looking at distributions and network patterns without any expectation or structure. In fact, this node implies a level of generality that is not amenable to sample size calculations. Nodes 2 and 3, in the middle of the workflow, rely on a choice of metrics for describing the microbiome distribution. In Node 2, the choice is made to describe the microbiome by betadiversity<sup>15</sup>, a measure of dissimilarity between samples. Sample size calculations are based on comparisons of dissimilarity within a group (i.e., within patients, within controls, or within samples following the same treatment) to distances between these groups<sup>16</sup>. In Node 3, the choice is made to represent the microbiome profile by the set of taxa abundances, which can then be compared between and within groups of samples. Nodes 4 to 7, at the bottom row of Fig. 1, refer to hypotheses that reduce microbiome data to a single number per sample. For example, within-sample richness can be characterized by alpha-diversity (Node 4). In Node 5, counts for one taxon of interest are considered; Node 6 focuses on whether a specific sample belongs to a particular cluster, such as a species subtype. In node 7, one might ask whether a sample contains species belonging to a specific taxon. Finally, after choosing the question of interest, sample size calculations would compare the effect size of the chosen measure between groups, see following sections for examples.

structure between groups versus within groups using betadiversity or distance metrics (Fig. 1, node 2)) describes how to setup sample size calculations for this situation, with links to recommended methods and formulae.

Simple calculations that assume well-behaved distances (e.g., normally distributed distances for pairs in the same group) can be performed using information that may be easy to extract from published papers. If the means and standard deviations of beta-diversity distances are reported, then sample size calculations based on Equation A (Table 2) can be used.

An analysis of variance (ANOVA) compares variability within a group to variability between groups, which is exactly the concept desired for beta-diversity analyses. Furthermore, when comparing only two groups, there is an algebraic equivalence between sample size calculations using the t-test (Equation A, Table 2), derived from an ANOVA F-test, or based on correlation (Equation G, Table 2), i.e., the square root of the model-captured  $R^2$  17; F-statistics or  $R^2$  values are often reported in publications, for example see Table 5 in Sugino et al. 18. Box 1 describes how the squared correlation,  $R^2$ , is related to the effect size  $\Delta$  from Equation A. It is worth noting that one should only use the  $R^2$  reported in a regression model for a microbiota beta-diversity sample size calculation when the publication has used the same distance measure as planned for one's own study.

Madan et al. compared beta diversity between infants who were born by vaginal delivery versus cesarean section<sup>19</sup>, and we use these data to illustrate calculations with Equation A. Means and standard deviations can be extracted by eye from Fig. 1b in their paper, and are shown in Table 4. Madan et al. did not provide estimates of  $R^{2 ext{ } 19}$ .

Using these values, we can estimate the sample size required to compare vaginal birth and cesarean section with 80% power. To perform conservative calculations, we assume a

### **Box 1.** Relationship between effect size, $\Delta$ , sums of squares from ANOVA, and correlation for studies comparing two groups

Within-group and between-group sums of squares (SSW, SSB, respectively) are basic elements of ANOVA; the ANOVA F-statistic is calculated as  $\frac{SSB/(G-1)}{SSW/(n-G)}$  where G is the number of groups and n is the sample size. Cohen<sup>17</sup> defined a statistic,  $f^2 = \frac{SSB}{SSW}$  where this formula can also be seen as the ratio of the variance between-group means to the variance of the data within the groups. If there are only two groups, then the effect size for a two-group t-test,  $\Delta$ , can be written  $\Delta = 2f$  where  $\Delta$  is the effect size from Equation A (Table 2). Furthermore, the  $R^2$  measure from a regression model or an ANOVA, can be written as:

$$R^2 = \frac{f^2}{(1 + f^2)}$$

when there are two groups. The square root of  $R^2$ , for the two-group comparison, is the  $\rho_{yx}$  of Equation G. For groups of unequal size, for more than two groups, or unequal standard deviations, adjustments to the simple formulas have been developed  $^{17}$ .

common standard deviation equal to the larger within-group value: 0.0046. Then, we can calculate the effect size from this study as  $\Delta = \frac{0.5613 - 0.5587}{0.0046} = 0.0026/0.0046 = 0.565$ . Thus, according to Equation A (Table 2) for a desired 5% type I error for two-sided testing, and 20% type II error, the required sample size, per group, to detect differences between vaginal birth and C-section is:

$$n_1\,=\,2\frac{(1.96\,+\,0.84)^2}{0.565^2}=\,49.11\,\approx\,50,$$

so, the total sample size needed for the two groups would be estimated as about 100.

We can also estimate  $f=\Delta/2=0.2825$ , and hence  $R^2=\frac{0.2825^2}{(1+0.2825^2)}=0.0739$ . Therefore, the correlation can be estimated as  $\hat{\rho}_{yx}=\sqrt{0.0739}=0.272$ . Then according to Equation G (Table 2),

Table 3. Key formulae or references for sample size and power calculations for the specific hypothesis being tested.

Section for examples and position in decision tree	Hypotheses for X – Y relationship	Equations, assumptions and/or references for equations
Measures of whole community composition		
Comparing microbial community structure between groups versus within groups using betadiversity or distance metrics, Fig. 1, node 2	<ul> <li>Outcome or Y variable is the beta-diversity between pairs of samples.</li> <li>Explanatory or X variable is group membership or a binary exposure measure.</li> </ul>	<ul> <li>Option 1. Normal distribution—Equation A (Table 2), comparing average distances or diversity measures between pairs from different groups versus pairs from the same group. This simple solution does not consider dependence between the pairs.</li> <li>Option 2. Distributions of Beta-diversity measurements are compared between pairs from the same group and pairs from different groups. These more complex methods explicitly use all the distances between and within groups. Example: Bray-Curtis distances<sup>16</sup>; summarized by<sup>47</sup>.</li> </ul>
Using an entire vector of abundances to describe the microbiome of a sample, Fig. 1, node 3	<ul> <li>Outcome or Y variable is a set of taxon abundances.</li> <li>Explanatory or X variable is group membership or a binary exposure measure</li> </ul>	■ Dirichlet-Multinomial method by Holmes et al. 46
Single-measure summary of community compo	osition, or a measure of one particular taxon	
Testing association between total microbial alpha-diversity, or taxon-specific alpha-diversity and an exposure or grouping variables, Fig. 1, node 4	<ul> <li>Outcome or Y variable is microbiome alpha-diversity.</li> <li>Explanatory or X variable is group membership or a binary exposure measure.</li> </ul>	<ul> <li>Normal distribution—Equation A (Table 2).</li> <li>Non-normal distribution—Equation A (Table 2) after estimating the mean and standard deviation from medians and quantiles<sup>49</sup> as described in Footnote 1.</li> </ul>
	<ul> <li>Outcome or Y variable is microbiome alpha-diversity</li> <li>Explanatory or X variable is a continuous exposure or vice versa.</li> </ul>	■ Equation G (Table 2) for correlations.
Testing association between taxon abundances and an exposure or grouping variable, Fig. 1, node 5	<ul> <li>Outcome or Y variable is the abundance of a particular microbial taxon or taxa (actual or transformed from median values—see Footnote 1), i.e., continuous measurements.</li> <li>Explanatory or X variable is group membership or a binary exposure measure.</li> </ul>	■ Equation A (Table 2) for mean abundances.
	<ul> <li>Outcome or Y variable is taxon abundance, represented by a binary variable where the abundances are categorized as high or low using a threshold such as the median.</li> <li>Explanatory or X variable is group membership or a binary exposure measure.</li> </ul>	<ul> <li>Equation C, or D (Table 2) for proportions, or Equation F (Table 2) for odds ratios.</li> <li>The calculations for odds ratios will be equivalent if Y and X are reversed.</li> </ul>
	<ul> <li>Outcome or Y variable is taxon abundance (a continuous quantity).</li> <li>Explanatory or X variable is a continuous exposure or vice versa.</li> </ul>	■ Equation G (Table 2) for correlations.
Testing higher or lower rates of cluster membership between groups, Fig. 1, node 6	<ul> <li>Outcome or Y variable is microbiota cluster membership, yes or no.</li> </ul>	■ Equation C or D (Table 2) for proportions, or Equation F (Table 2) for odd ratios.
	<ul> <li>Explanatory or X variable is group membership or a binary exposure measure or vice versa.</li> </ul>	■ The calculations for odds ratios will be equivalent if Y and X are reversed.
Testing higher or lower rates of taxon membership between groups (i.e., colonization with a microbe), Fig. 1, node 7	<ul> <li>Outcome or Y variable is taxon membership (colonization with a microbe), yes/no.</li> <li>Explanatory or X variable is group membership or a binary exposure measure.</li> </ul>	<ul> <li>Equation C or D (Table 2) for proportions, or Equation F (Table 2) for odd ratios.</li> <li>The calculations for odds ratios will be equivalent if Y and X are reversed.</li> </ul>
	measure.	

the required sample size for testing 
$$\rho_{yx} = 0$$
 (assuming  $h = 0$ ) is: 
$$n = 3 + \frac{(1.96 + 0.84)^2}{(\log_e((1 + 0.272)/(1 - 0.272))/2)^2} = 103.71 \approx 104$$

This is slightly larger than the results based on Equation A provided above. These two estimates agree very well for larger sample sizes, but Equation G tends to be more conservative for small sample sizes<sup>20</sup>.

We could also estimate sample size using estimates of SSB and SSW calculated from the same information in Table 4. Following the

Table 3. continued

Section for examples and position in decision tree	Hypotheses for $\mathbf{X} - \mathbf{Y}$ relationship	Equations, assumptions and/or references for equations
Microbiome, exposure, and phenotype		
Microbiome as the mediator (exposure-microbiome-outcome)	■ Microbiome is the mediator $M$ (exposure $X \rightarrow$ microbiome $M \rightarrow$ outcome $Y$ ).	■ To calculate the required sample sizes to detect mediation in a study, one must
		obtain estimates of $\alpha$ (alpha, correlation between $X \& M$ , i.e., between an exposure and microbiome) and $\beta$ (beta, correlation between $X$ and $Y$ , i.e., between microbiome and outcome) from preliminary work or the literature. The required sample size will depend on the mediation test chosen. Supporting information for samples sizes can be found in Table 3 of Fritz and MacKinnon <sup>35</sup> .

For total sample size ≥25, the median itself is the best estimator of the mean. For total sample size <25, then the following formula can be used to estimate the mean:  $mean \approx \frac{a+2m+b}{4}$ , where m= the median, a is the minimum value, and b is the maximum value<sup>26</sup>. For total small sample size <15, the standard

deviation (square root of the variance) can be estimated using the following formula:  $SD = \sqrt{\frac{1}{12}(\frac{(a-2m+b)^2}{4}+(b-a)^2)}$ . For moderately sized samples, the range/4 is the best estimator for the standard deviation. For large samples (size more than 70), Range/6 provides a better estimate for the standard deviation<sup>49</sup>. If only an interquartile range is provided, the standard deviation can be estimated by the range divided by 1.35<sup>50</sup>, an expectation built on the expected quantiles of the normal distribution.

principles in Box 1, the sum of squares for the vaginal birth group can be calculated as  $70 \times 0.0026^2 = 4.732 \times 10^{-4}$ , and the sum of squares for the cesarean group is  $32 \times 0.0046^2 = 6.771 \times 10^{-4}$ . Therefore, SSW is their sum, i.e., 0.001150. To calculate SSB, the overall mean is first obtained by a weighted average, as

$$(0.5613 * 70 + 0.5587 * 32)/(70 + 32) = 0.56048.$$

Therefore, SSB is  $70 * (0.5613 - 0.56048)^2 + 32 * (0.5587 (0.56048)^2 = 0.000148$ 

With SSW and SSB in hand, then:

$$f^2 = \frac{SSB}{SSW} = 0.000148/0.001150 = 0.1287, f = \sqrt{0.1287} = 0.3587.$$

Hence  $R^2 = \frac{0.3585^2}{(1+0.3585^2)} = 0.1140$ . Therefore, the correlation can be estimated as  $\rho^*_{yx} = \frac{1}{2} \left( \frac{1}{2} + \frac{1}{$  $\sqrt{0.1140} = 0.3376$ , and according to Equation G (Table 2), the required sample size for testing  $\rho_{vx} = 0$  is:

$$n = 3 + \frac{(1.96 + 0.84)^2}{(\log_e((1 + 0.3374691)/(1 - 0.3374691))/2)^2}$$
  
= 66.55 \approx 67

This sample size estimate of 67 is based on the two standard deviations shown in Table 4, whereas the earlier calculation based on Equation A used the larger of the two, which explains the discrepancy between the two sample size estimates.

These simple beta-diversity sample size calculation are based on normality of the within-group pair distances. A richer approach that relaxes this assumption can be built on the full distribution of pairwise distances, then analyzing the data using concepts from multivariate statistics. Since distance distributions tend to be strongly skewed, such alternative methods are commonly used for analysis after collecting study data. However, performing sample size calculations for these nonparametric analyses is challenging, since information about the distribution of the distances is needed. In a previous paper by the IMPACTT consortium we described approaches for estimating sample size and power when distances are available<sup>21</sup>. Furthermore, since these distances are often difficult to obtain, we also demonstrated how to generate distances by simulation<sup>21</sup>.

### Using an entire vector of abundances to describe the microbiome of a sample (Fig. 1, node 3)

Multivariate methods can be used to compare microbial community structures through examination of distributions of the counts of taxa abundances. These distributions tend to have a large and heavily skewed dynamic range, with some very large counts and many near zero. No simple distributions match the shape and variability well, and hence specific methods for sample size calculation are needed. Row Using an entire vector of abundances to describe the microbiome of a sample (Fig. 1, node 3) in Table 3 shows the usual setup and method. Although resampling-based comparisons could be considered (e.g., permutation tests), they rest on assumptions which may not hold, such as that the within-group variability is consistent across groups. Therefore, La Rosa et al.<sup>22</sup> proposed tests for comparing community structures based on the Dirichlet-Multinomial distribution. Since their approach is based on parametric distributions, it contains parameters that can be interpreted as measures of how different the community structures are. Thus, their method can be expected to be more powerful than any nonparametric procedure. The combination of the Dirichlet distribution with the multinomial allows capture of the intersample variability needed for microbiome data; in statistics this feature is referred to as 'over-dispersion'. There are two key parameters:  $\pi = (\pi_1, ..., \pi_k)$  represents the expected taxa frequencies averaged across the groups being compared, and  $\theta$ represents the over-dispersion.

In La Rosa et al.<sup>22</sup> three tests are introduced and demonstrated: (a) comparing one community structure to an expectation, i.e.,  $H_0: \pi = \pi_0$ , where  $\pi_0$  is known, (b) comparing two groups,  $H_0: \pi_1 = \pi_2$ , and (c) comparing multiple groups,  $H_0: \pi_1 = \pi_2 = \pi_3 = \dots$  All these tests, and corresponding power calculations are built into their software package:

To give an example, here we describe their calculations comparing community structures between two groups. Their

**Table 4.** Means and standard deviations of Unifrac beta diversity among 6-week-old infants by mode of delivery, extracted from Fig. 1b of Madan et al.<sup>19</sup>.

	Vaginal	Cesarean section
n per group	70	32
Mean	0.5613	0.5587
Standard deviation	0.0026	0.0046

data were taken from three oral sites (subgingival, supragingival, and saliva) in 24 subjects of both genders from the USA. Power calculations are based on a modified version of Cramer's  $\varphi_{\square}$  criterion,  $\varphi_m$ , which is based on a contingency table chi-squared test statistic ( $\chi^2$ ):

$$\varphi_m = \sqrt[2]{\frac{\chi^2}{\chi^2_{\max}}}$$

The value of this normalized chi-squared statistic is determined by the two key parameters,  $\pi$  and  $\theta$ . When the authors compared the distributions for subgingival plaque and supragingival plaque in their subjects, their value of the modified Cramer's  $\varphi_m$  was 0.16. They then calculated power for different numbers of subjects per group, and for different numbers of reads, at significance thresholds of 1 and 5%. For 1000 reads per group, power increased from 29.46% with 10 subjects per group to 89.76% with 25 subjects per group, using a significance threshold of 1%. It is worth noting that the authors recommend aggregating very rare taxa with abundance <1% into a single category.

## Testing association between total microbial alpha-diversity, or taxon-specific alpha-diversity and an exposure or grouping variables (Fig. 1, node 4)

In community ecology, alpha-diversity refers to the number of species present in an ecosystem (richness)<sup>15</sup> as well as the frequency of occurrence of each type of organism (evenness). This ecological metric is found to be reduced in several disease states<sup>24</sup>, making it a relevant factor to consider when proposing microbiome research hypotheses. The most commonly used metrics/indices are Shannon, Inverse Simpson, Simpson and Chao indices<sup>25</sup>. These indices do not consider the phylogeny of the taxa identified in sequencing. One measure of phylogenetic diversity (Faith's PD) is based on phylogeny and can be calculated when a microbial phylogenetic tree is available<sup>24</sup>.

When alpha-diversity is normally distributed or can be log-transformed, basic equations can be used for sample size calculations—see row Testing association between total microbial alpha-diversity, or taxon-specific alphadiversity and an exposure or grouping variables (Fig. 1, node 4), node 4 in Table 3.

For an example of how to calculate sample size with an alphadiversity metric, we will consider a study presented by Casals-Pascual et al. which used Faith's phylogenetic diversity (Faith's PD). This study aimed to compare the diversity of gut microbial communities in two phenotypically distinct groups of patients with Crohn's disease (CD). The null hypothesis was that gut microbiota phylogenetic diversity did not differ by CD phenotype. To test this hypothesis, CD patients with the B1 phenotype would be compared with those with either a B2 or B3 phenotype. In this case, the CD phenotype was the independent variable and gut microbial diversity (Faith's PD), the dependent variable.

To determine the number of patients required to find a statistically significant difference in Faith's PD between CD phenotypes, researchers searched for summary statistics on Faith's PD. They found a gut microbiota study of 100 patients with the B1 CD phenotype that reported a standard deviation of 3.45 for Faith's PD, a mean of 13.5, and the distribution seemed to be

approximately normal. To determine a clinically meaningful effect size, and because a similar previous study did not exist, researchers considered an analogous study where patients treated with antibiotics were compared to healthy controls. In the analogy, an effect size of 1.5 units was observed with Faith's PD metric with a significance level of 0.0001. Using Equation A in Table 2 and a standard deviation of 3.45, selecting a conventional level of statistical significance of 5% and a statistical power of 80%, a total sample size of 110 patients (55 per group) was recommended to detect differences in Faith's PD of ≥2 units<sup>7</sup>.

The median value can also be used to calculate sample size and is particularly appropriate for skewed richness and diversity values; the formulae needed for converting medians and interquartile ranges to means and standard deviations are shown and referenced in Footnote 1 Table 3.

When the exposure variable is continuous, methods based on correlations can be used. For example, soil bacteria metagenome alpha-diversity has been associated with mean annual precipitation gradients<sup>26</sup>. Suppose a researcher wants to test the null hypothesis that alpha-diversity is unrelated (r=0) to mean annual precipitation with  $\alpha=0.05$  and power of 0.95. The researcher assumes that the alternative correlation coefficient (h) is approximately -0.5. Therefore, following Equation G in Table 2:

$$\tilde{\rho}_{yx}^* = \frac{1}{2} \ln \left( \frac{1 + \tilde{\rho}_{yx}}{1 - \tilde{\rho}_{yx}} \right) = \frac{1}{2} \ln \left( \frac{1 + (-0.5)}{1 - (-0.5)} \right)$$

$$= -0.549$$

$$z_{1-\frac{\alpha}{2}} = 1.96(\alpha = 0.05); z_{1-\beta} = 1.64(\beta = 0.05)$$

For a null hypothesis ( $H_0$ ) of no correlation, the required sample size is approximately:

$$n = 3 + \left( \left( z_{\left(1 - \frac{\alpha}{2}\right)} + z_{1 - \beta} \right)^{2} / \left( \tilde{\rho}_{yx}^{*} - h^{*} \right)^{2} \right)$$

$$= 3 + \left( (1.96 + 1.64)^{2} / (-0.549 - 0)^{2} \right)$$

$$= 45.9 \approx 46$$

### Testing association between taxon abundances and an exposure or grouping variable (Fig. 1, node 5)

Researchers may also hypothesize about abundances of a specific microbial taxon of interest. These hypotheses can be expressed either using mean abundances, or by examining the proportion of samples with abundances over a chosen threshold. Sample size calculations can be based on either choice of metric, see row *Testing association between taxon abundances and an exposure or grouping variable* (Fig. 1, node 5) corresponding to node 5 in Table 3.

An important consideration here, as well as for nodes 6 and 7 in Fig. 1, is the choice of type 1 error. If only one taxon is of interest, then the type 1 error threshold,  $\alpha$ , should not need adjustments for multiple testing. However, if the study plans to test association at all available taxons, then power calculations should be performed using a value  $\alpha^*$  which controls family-wise error rate. For example, use of the Bonferroni correction would suggest  $\alpha^* = \alpha/M$  where M is the planned number of tests.

For example, in the observational study of Koleva et al.<sup>27</sup> looking at the gut microbiome of mother-infant pairs (total population size of 1,021) from the Canadian Healthy Infant Longitudinal Development (CHILD) Study, the authors hypothesized the genus *Lactobacillus* was reduced in gut microbiota of male infants born to an asthmatic mother<sup>27</sup>. The abundance of 16S data from fecal samples collected at 3–4 months after birth was compared between infants born to mothers who received asthma treatment during pregnancy (i.e., infants high risk for

**Table 5.** Percent distribution of highest *Lactobacillus* abundance (highest tertile cut-off) between asthmatic mothers and control group (adapted from Table 1 of Koleva et al.<sup>27</sup>.

Group	Highest <i>Lactobacillus</i> abundance		Total
	Yes	No	
Asthma	9 (10%)	78 (90%)	87
No asthma	246 (26%)	688 (74%)	934

**Table 6.** Median relative abundance (interquartile range) of Bacteroidetes in female infants of mothers with and without asthma (Adapted from Table 3 of Koleva et al.<sup>27</sup>.

Group	Bacteroidetes
Asthma ( $n_1 = 17$ )	72.8 [21.8–80.2]
No asthma $(n_2 = 145)$	31 [0.5–62.3]

allergic diseases) and those who were not. Results supported the *Lactobacillus* hypothesis in a sex-dependent and ethnicity-dependent manner<sup>27</sup>. In male Caucasian infants, the reduction of *Lactobacillus* was independent of other study covariates known to also influence the infant gut microbiome, such as prepregnancy overweight, atopy status, breastfeeding and intrapartum antibiotic treatment, strengthening these conclusions. Infant fecal *Lactobacillus* abundance was transformed into a binary variable using the cut-off value for the highest tertile (Table 5).

The R script in Box 1 of the Supplement can be used to implement the sample size and power calculations using equations D and E of Table 2. However, we illustrate calculations assuming the sample sizes are equal in each group (r=1, Equation C of Table 2). Assuming a 2-sided test with an  $\alpha$  of 0.05, 87 samples in each group can provide 80% power to detect a difference of this size in the proportion of infants with Lactobacilli above the highest tertile.

In the same study, abundances of bacterial taxa other than *Lactobacillus* were made using the Benjamini–Hochberg method<sup>28</sup> to adjust for multiple testing (which is built into the multi-test procedure in SAS). Tests for interactions between infant sex and maternal prenatal asthma on *Lactobacillus* abundance were performed using an adjusted rank transform (ART) nonparametric test.

The taxon abundance comparisons presented in this paper are also useful to calculate sample size for other research questions. Due to the non-normal distribution of taxon abundance data, we provide an example that first requires converting median abundance into mean abundance for use in the sample size equation (Table 6). This conversion may be more appropriate at the phylum or other higher classification level, even family level, in which abundance data may be least skewed. Nevertheless, for illustration we used the median abundance of fecal Bacteroidetes in female infants of mothers with and without asthma to calculate sample size (Table 6). Medians and IQR (Q1 and Q3) were provided in the paper, and we transform these to estimate the mean and standard deviation. In this case, we use mean = median and SD = IQR/1.35:

Mean = Median 
$$SD = \frac{IQR}{1.35} = \frac{Q_3 - Q_1}{1.35}$$

Therefore, in female infants of mothers with asthma, the mean and SD are estimated to be:

Mean = median = 72.8; 
$$n_1 = 17$$
;  $SD_1 = \frac{Q_3 - Q_1}{1.35} = \frac{80.2 - 21.8}{1.35} = 43.26$ 

In female infants of mothers without asthma, the mean and SD will be:

Mean = median = 
$$31; n_1 = 145; SD_1 = \frac{Q_3 - Q_1}{1.35} = \frac{62.3 - 0.5}{1.35} = 45.78$$

$$SD_{pooled} = \sqrt{\frac{(n_1 - 1)SD_1^2 + (n_2 - 1)SD_2^2}{n_1 + n_2 - 2}}$$
$$= \sqrt{\frac{(17 - 1)*43.26^2 + (145 - 1)*45.78^2}{17 + 145 - 2}} = 45.53$$

The effect size,  $\Delta$ , is therefore:

$$\Delta = \frac{\textit{M}_1 - \textit{M}_2}{\textit{SD}_{pooled}} = \frac{72.8 - 31}{45.53} = 0.918.$$

Based on this effect size, we calculate sample size using Equation A in Table 2 for testing differences in means:

$$n=\frac{2\left(z_{1-\frac{\alpha}{2}}+z_{1-\beta}\right)^{2}}{\Delta^{2}},$$

where, n is the sample size in each group (assuming sizes of the two groups are equal),  $z_{1-\frac{\alpha}{2}}=1.96$  ( $\alpha=0.05$ ), and  $z_{1-\beta}=0.84$  ( $\beta=0.20$ ). Therefore:

$$n = \frac{2(z_{1-\frac{\alpha}{2}} + z_{1-\beta})^2}{\Delta^2} = \frac{2 * (1.96 + 0.84)^2}{0.918^2} = 18.60 \approx 19$$

Assuming a 2-sided  $\alpha$  of 0.05, 19 samples in each group can achieve 80% power to compare Bacteroidetes abundance between asthmatic and non-asthmatic mothers of female infants.

### Testing higher or lower rates of cluster membership between groups (Fig. 1, node 6)

Microbiota community types or clusters are increasingly being used to characterize whole microbial community composition. The format of these variables is categorical, and therefore sample size calculations are straightforward, see row *Testing higher or lower rates of cluster membership between groups* (Fig. 1, node 6) for node 6 in Table 3.

We provide an example here based on a study by Tun et al. (2021)<sup>29</sup> that identified 4 longitudinal gut microbiota clusters during infancy (Table 7). A sample size calculation is performed to determine the association between Asian ethnicity vs. any other, and the presence or absence of the C1–C1 cluster in the infant. A binary variable was created for the presence of the C1–C1 cluster vs. three other clusters.

R code and results for estimating sample size and power from these data, following Equations C, E in Table 2, are shown in the Supplementary Material, Box 2. There, we consider the sample sizes are equal in each group (r=1). Assuming a 2-sided test with an  $\alpha$  of 0.05, 138 samples in each group can achieve 80% power to find an association between Asian ethnicity (vs. other) and the C1–C1 gut microbiota cluster (vs. others).

### Testing higher or lower rates of taxon membership between groups (i.e., colonization with a microbe) (Fig. 1, node 7)

To determine sample sizes for colonization with specific microbiota (presence/absence or yes/no), one can use the equations for proportions or odds ratios (Table 3, row Testing higher or lower rates of taxon membership between groups (i.e., colonization with a microbe) (Fig. 1, node 7) for node 7).

To illustrate sample sizes for colonization with specific microbiota (presence/absence or yes/no), we draw information from Drall et al.<sup>30</sup> to test the question whether infant *C. difficile* colonization differs by exclusivity of breastfeeding (Table 8). In 853 exclusively breastfed infants (EBF), the *C. difficile* colonization rate

**Table 7.** Percent distribution of the C1–C1 cluster vs. the other three clusters between Asian ethnicity and others (Table 1 of Tun et al. (2021))<sup>29</sup>.

Group	Gut microbiota cluster		Total
	C1-C1	Others	
Asian ethnicity	51 (30%)	120 (70%)	171
Others	23(16%)	123 (84%)	146

**Table 8.** *C. difficile* colonization rate between exclusively breastfed infants vs. partial breastfed or formula-fed infants.

Mode of feeding	C. difficile colonization
Exclusively breastfed infants	22.63%
Others	41.23%

was 22.63%, i.e., 193 exclusively breastfed infants were colonized with *C. difficile*. In 431 partially breastfed infants (PBF), the *C. difficile* colonization rate was 35.96%; hence 155 partially breastfed infants were colonized, and in 270 exclusively formula-fed infants (EFF), the *C. difficile* colonization rate was 49.63% implying 134 colonized infants (Table 8).

The pooled proportion of *C. difficile* colonization from the PBF and EFF groups is obtained as:

$$P = \frac{\textit{No. colonized cases in PBF} + \textit{EFF}}{\textit{Number of infants in PBF} + \textit{EFF}} = \frac{155 \, + \, 134}{431 \, + \, 270} = \, 0.4123$$

R code and results for estimating sample size and power from these data, assuming an equal sample size in each of the two groups (r=1), can be found in Supplementary Material, Box 3. Calculations follow Equation C in Table 2. Assuming a 2-sided test with an  $\alpha$  of 0.05, 95 samples in each group are needed to achieve 80% power to find differences in *C. difficile* colonization rates between exclusively breastfed infants and other infants.

This sample size calculation example will apply to any comparison of populations with respect to presence or absence of a microbe of interest. For example, this approach would be appropriate for a study comparing the presence of shared microbial species between animals and humans. The sample size calculation might also be used to compare samples where the entry or exit of microbial species into/from an ecosystem is expected. For instance, a study where a single species (probiotic) or an entire community is introduced, such as fecal microbiota transplantation.

### MICROBIOME AS THE MEDIATOR (EXPOSURE-MICROBIOME-OUTCOME)

A mediator variable (M) explains part or all of the relationship between an independent variable (X) and a dependent variable (Y), and the question of whether microbiome mediated relationships between exposures and disease is highly topical. One common focused research question is whether all association between X and Y passes through X, i.e., complete mediation (see the row Microbiome as the mediator (exposure-microbiome-outcome), Table 3).

Approaches to test for mediation vary with the most common approach being the Baron and Kenny's Causal-Steps test<sup>31</sup>. The four steps of this approach are: (i) the total effect of X on Y must be present (statistically significant), (ii) there must be an effect of X on M, (iii) M must have a non-zero effect on Y even after controlling for X, and (iv) the effect of X on Y controlling for X must be smaller than the total effect of X on Y. All four criteria must be satisfied to consider mediation through X to be

present. Other mediation tests include the Joint Significance Test<sup>32</sup>, which is a variation of Baron and Kenny's test, and the product-of-coefficients tests, Sobel First-Order<sup>33</sup> and PROD-CLIN<sup>34</sup> tests. Valuable information on these tests and how they compare to each other has been summarized by Fritz and MacKinnon<sup>35</sup>.

For instance, to determine the sample size to assess whether infant gut microbiota could be in the biological pathway from maternal prenatal overweight to offspring overweight, we refer to the correlation coefficients in Fig. 3b of Tun et al. 36 that tested the mediating effect of the Lachnospiraceae in infant gut microbiota on the association between maternal pre-pregnancy overweight and child overweight. The values for the correlations presented in Tun et al.<sup>36</sup> mediation Fig. 3b could be obtained separately from other studies that did not pursue mediation analyses. Their Fig. 3b indicates that the correlation between pre-pregnancy overweight and fecal Lachnospiraceae abundance is small (r = 0.11), and that the correlation between fecal Lachnospiraceae abundance and child overweight is of medium size (r = 0.41). Looking up these values in Table 3 of<sup>35</sup>, we determined the required total sample size to be 400-427 depending on the mediation method chosen.

### SPECIAL CONSIDERATIONS FOR LONGITUDINAL STUDIES AND OTHER STUDY DESIGNS

Due to the dynamic nature of the microbiome, longitudinal microbiome data are usually more informative about the profile of microbiome in relation to its host and environmental interactions<sup>37</sup>. Subjects are measured repeatedly during a study, which allows for direct evaluation of changes of response variable over time, as well as more precise estimates of inter-individual differences. Hence, longitudinal studies allow for both between-subject differences and within-subject dynamics to be considered, resulting in more powerful studies<sup>38,39</sup>.

Current efforts in microbiome research aim to move from correlation to causality. Longitudinal studies are invaluable research tools to evaluate a causative impact of the microbiome on the host physiological disease processes. Throughout the lifetime, our microbiome changes constantly over body habitats and time, it can be completely modified, either temporarily or permanently, by diseases such as infection, or medical interventions such as antibiotic courses<sup>40</sup>. These temporal patterns can help reveal if changes in microbiome predict, create or prevent diseases<sup>41</sup>. Causal inference analysis methods, such as mediation analysis, are recommended in longitudinal studies to evaluate the mutual relationship between the microbiome, the host and other study variables (environment, intervention, etc.). Various computational methods such as regression-based time series models, autoregressive (AR) models, and richer models such as the microbiome counts trajectories infinite mixture model<sup>37</sup> have been applied to longitudinal microbiome data. However, there are still methodological limitations to existing approaches when coping with current methods of dynamic and complex microbiome data.

While longitudinal studies are more powerful biologically and statistically, calculating sample size for a longitudinal study with repeated microbiome measures is complex, and requires simulation methods based on pilot datasets <sup>42</sup>. Ideally, to run this type of simulation, the pilot dataset should contain the independent (microbiome) and dependent (host factor) variables measured at multiple timepoints, preferably the same timepoints planned for the upcoming study. It is essential to consider the between-patient covariance in both the dependent and independent measures, as well as within-patient covariance between time-points. It is nearly impossible to accurately guess or simulate these covariance structures, and to obtain reproducible sample size calculations without having longitudinal pilot data.

#### **CONCLUDING REMARKS**

The goal of this resource is to provide a framework to help plan sample size calculations in microbiome studies. We have illustrated several options such as a decision tree in Fig. 1, and provided various study designs, formulas, as well as worked examples in Sections "Sample size calculations associated with each node of the decision tree and Microbiome as the mediator (exposure-microbiome-outcome)". We also provided codes to implement equations in R in Supplementary Material. It is often not straightforward to test associations between potential environmental factors or phenotypes and microbiome composition, whether measured by OTUs, ASVs or taxa abundances. Microbiome data often display a broad dynamic range, high dimensionality, substantial variability in counts between samples, and non-normality of counts. Counts tend to be correlated with each other, possibly due to phylogenetic structure or similarity of function<sup>43</sup>. Furthermore, diversity between samples can be extreme, such that some species or OTUs are absent from some samples. Statistically, these characteristics must be considered to perform valid analysis and sample size calculations. Evaluating the sample size threshold to design meaningful microbiome studies remains a critical step. By expanding from previously published methods, the approaches presented here will help design interpretable human microbiome studies. We encourage other scientists to test and optimize the tools presented here, and future ones, as microbiome datasets increase in public repositories.

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#### **COMPETING INTERESTS**

The authors declare no competing interests.

#### **ADDITIONAL INFORMATION**

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